

Filtering with limited information

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Lars Peter Hansen (2014)

Hansen (1982) builds on a long tradition in econometrics of 'doing something without having to do everything.' This entails the study of partially specified models, that is, models in which only a subset of economic relations are formally delineated.

- Filtering shocks from a dynamic model is a central task in macroeconomics:
 - 1. Path of shocks is a reality check for the model.
 - 2. Historical decompositions and counterfactuals.
 - 3. Forecasting and optimal policy recommendations when laws of motions are state-dependent.
 - 4. Structural estimation.
- However, filtering is hard, and there is no universal and easy-to-apply algorithm to implement it.

Alternatives

- A possibility ⇒ sequential Monte Carlos: Fernández-Villaverde, Rubio-Ramírez, and Schorfheide (2016).
 - 1. Specification of the full model, including auxiliary assumptions.
 - 2. Computationally costly.
 - 3. Curse of dimensionality.
- Some of the previous points also hold for even the simple linear, Gaussian case where we can apply the Kalman filter.
- Can we follow Hansen's suggestion and 'do something without having to do everything"? ⇒ Yes!
- Partial information filter.

• Model:

$f(x_t, y_t, \mathbb{E}_t[g(x_{t+1}, y_{t+1}, x_t, y_t)]) = \mathbf{0}$

• Deterministic steady state:

 $f(\bar{x},\bar{y},g(\bar{x},\bar{y},\bar{x},\bar{y})])=\mathbf{0}$

• Estimation:

$$f(\widehat{x}_t, y_t, \widehat{\mathbb{E}}_t[g(\widehat{x}_{t+1}, y_{t+1}, \widehat{x}_t, y_t)]) = \mathbf{0}$$

Factorization

• Factorization of $g(\circ)$

 $g(x_{t+1}, y_{t+1}, x_t, y_t) \equiv g_1(x_{t+1}, y_{t+1}, x_t, y_t) \times g_2(x_{t+1}, y_{t+1}, x_t, y_t)$

• Then:

 $\mathbb{E}_t[g(x_{t+1}, y_{t+1}, x_t, y_t)_i] \equiv \mathbb{E}_t[g_1(x_{t+1}, y_{t+1}, x_t, y_t)_i] \times \mathbb{E}_t[g_2(x_{t+1}, y_{t+1}, x_t, y_t)_i] \\ + \operatorname{Cov}_t[g_1(x_{t+1}, y_{t+1}, x_t, y_t)_i, g_2(x_{t+1}, y_{t+1}, x_t, y_t)_i]$

- We need to approximate conditional first and second moments:
 - 1. Equilibrium conditions of the model.
 - 2. Observed expectations.
 - 3. Auxiliary statistical model.

Auxiliary statistical model

• VAR(1) in $g_{1,t}, g_{2,t}$, (a subset of $n_{\tilde{y}}$ elements of) y_t , and \hat{x}_t for t = 1, ..., T.

• Why?

• If we collect variables collected in ξ_t :

$$\xi_t = \mu + A\xi_{t-1} + \epsilon_t, \quad Var[\epsilon_t] = \Sigma$$

• Ordering $g_{1,t}$ and $g_{2,t}$ as the first two variables of the VAR, we can write:

 $\mathbb{E}_t[g(x_{t+1}, y_{t+1}, x_t, y_t)]_i \equiv (\mu_i + e'_i A \xi_t)(\mu_{i+m} + e'_{i+m} A \xi_t) + \Sigma_{i,i+m}$ where e_i is a selection vector.

• Thus:

$$f(x_t, y_t, (\mu_i + e'_i A \xi_t)(\mu_{i+m} + e'_{i+m} A \xi_t) + \Sigma_{i,i+m})_i = 0$$

Two approaches

• (Fixed point): we find \hat{x}^{T} (and associated parameters for the VAR) that solve for all *i*:

 $f(x_t, y_t, (\mu_i + e'_i A \xi_t)(\mu_{i+m} + e'_{i+m} A \xi_t) + \Sigma_{i,i+m})_i = 0$

In practice, initialize $\widehat{E}_t^{(0)}[g(\circ)]$ based on $\widehat{x}_t^{(0)} = \overline{x} \forall t$ and $A^{(0)} = \mathbf{0}$, $\mu^{(0)} = \overline{\xi}_T$ and $\Sigma^{(0)} = \mathbf{0}$ and iterate until convergence.

- (Gibbs sampler): From d = 1, ..., D, iterate on:
 - 1. Given $\{\hat{x}_{t}^{(d)}\}_{t=1}^{T}$, $\mu^{(d)}, A^{(d)}, \Sigma^{(d)} \sim P(\mu, A, \Sigma | \xi_{t})$.
 - 2. Given $y^T, \mu^{(d)}, A^{(d)}, \Sigma^{(d)}$, solve for $\{\hat{x}_t^{(d+1)}\}_{t=1}^T$ for all i in:

 $f(x_t, y_t, (\mu_i + e'_i A \xi_t)(\mu_{i+m} + e'_{i+m} A \xi_t) + \Sigma_{i,i+m})_i = 0$

Start the Gibbs sampler from $\hat{x}_t^{(0)} = \bar{x}$ or the fixed point above. The Gibbs sampler allows us to quantify estimation uncertainty.

Tobin's Q

• Representative household

$$\max_{\{c_{t}, n_{t}, i_{t}\}} \mathbb{E} \sum_{s=0}^{\infty} \left(\prod_{u=1}^{s} \beta_{t-1+s} \right) u(c_{t+s}, n_{t+s})$$

s.t.
$$k_{t+1} = (1 - \delta_{t})k_{t} + e^{\xi_{t}} \left(1 - \frac{\chi}{2} \left(\frac{i_{t}}{i_{t-1}} - (1 + g) \right)^{2} \right) i_{t}$$

$$c_{t} + i_{t} = e^{z_{t}} k_{t-1}^{\alpha} n_{t}^{1-\alpha}$$

• Four shocks: $\beta_t, \delta_t, \xi_t, z_t$ follow log-linear AR(1) processes

Solution (I)

$$c_t^{-\eta}(1-\kappa(1-\eta)n_t^{1+1/\phi})^{1-\eta}=\lambda_t$$

$$c_t^{1-\eta}(1-\kappa(1-\eta)n_t^{1+1/\phi})^{-\eta}(1+1/\phi)\kappa(1-\eta)n_t^{1/\phi} = \lambda_t(1-\alpha)e^{z_t}\left(\frac{k_{t-1}}{n_t}\right)^{\alpha}$$

$$\lambda_{t} = \mu_{t} e^{\xi_{t}} \left(1 - \frac{\chi}{2} \left(\frac{i_{t}}{i_{t-1}} - (1+g) \right)^{2} - \chi \left(\frac{i_{t}}{i_{t-1}} - (1+g) \right) \frac{i_{t}}{i_{t-1}} \right)$$

$$+\beta_t \mathbb{E}_t \left[\mu_{t+1} \chi \left(\frac{i_{t+1}}{i_t} - (1+g) \right) \left(\frac{i_{t+1}}{i_t} \right)^2 \right]$$

$$\mu_t = \beta_t \mathbb{E}_t \left[(1 - \delta_{t+1}) \mu_{t+1} + \lambda_{t+1} \alpha e^{z_{t+1}} \left(\frac{n_{t+1}}{k_t} \right)^{1-\alpha} \right]$$

Solution (II)

$$c_t^{-\eta}(1-\kappa(1-\eta)n_t^{1+1/\phi})^{1-\eta} = \lambda_t$$

$$c_t^{1-\eta}(1-\kappa(1-\eta)n_t^{1+1/\phi})^{-\eta}(1+1/\phi)\kappa(1-\eta)n_t^{1/\phi} = \lambda_t(1-\alpha)e^{z_t}\left(\frac{k_{t-1}}{n_t}\right)^{\alpha}$$

$$1 = q_t e^{\xi_t} \left(1 - \frac{\chi}{2} \left(\frac{i_t}{i_{t-1}} - (1+g) \right)^2 - \chi \left(\frac{i_t}{i_{t-1}} - (1+g) \right) \frac{i_t}{i_{t-1}} \right)$$

$$+\mathbb{E}_t \left[\beta_t \frac{\lambda_{t+1}}{\lambda_t} q_{t+1} \chi \left(\frac{i_{t+1}}{i_t} - (1+g)\right) \left(\frac{i_{t+1}}{i_t}\right)^2\right]$$

$$q_t = \mathbb{E}_t \left[\beta_t \frac{\lambda_{t+1}}{\lambda_t} \left((1 - \delta_{t+1}) q_{t+1} + r_{t+1}^k \right) \right]$$

where $r_t^k \equiv \alpha e^{z_{t+1}} \left(\frac{n_{t+1}}{k_t} \right)^{1-\alpha}$

Quantitative set-up

- Calibration:
 - 1. $\bar{n} = \frac{1}{3}$.
 - 2. $\bar{g} = 0.5\%$, a 2% annual growth rate.
 - 3. $\eta = 1.0$, to have separable preferences for now.
 - 4. $\phi = 1$, a typical value for the Frisch elasticity.
 - 5. $\overline{\delta} = 2\%$, implying an 8% annual depreciation rate.
 - 6. $\beta = 0.995$, implying a annualized real rate of about $2\% + 4\eta \bar{g} = 5\%$.

7.
$$\rho_z = 0.95, \ \sigma_z = 0.76\%$$

- 8. $\rho_{\beta} = 0.8, \ \sigma_{\beta} = 1.0\%.$
- 9. $\rho_{\xi} = 0.9, \ \sigma_{\xi} = 1.0\%.$
- 10. $\rho_{\delta} = 0.75, \ \sigma_{\delta} = 1.0\%.$
- Solved the model using 3rd order perturbation methods with pruning, as in Andreasen, Fernández-Villaverde, and Rubio Ramírez (2017).
- Simulated data: We simulate the model for 2,000 periods after a burn-in of 1,000 periods.

Filter set up

- We want to filter a single variable, q_t using data on the SDF $M_{t+1} \equiv \beta_t \frac{\lambda_{t+1}}{\lambda_t}$, the rental rate on capital r_t^k , and the risk-free rate r_t^f .
- Set $x_t = q_t$, $y_t = [\tilde{M}_t, r_t^k, r_t^f]$, and rewrite the conditional expectation of the return on investment as:

$$g(x_{t+1}, y_{t+1}, x_t, y_t) \equiv y_{1,t+1} \times \left((1 - \bar{\delta}) x_{t+1} + y_{2,t+1} \right)$$
$$= \widetilde{M}_{t+1} \times \left((1 - \bar{\delta}) q_{t+1} + r_{t+1}^k \right)$$

with $\widetilde{M}_{t+1} = \overline{\beta} \frac{c_t}{c_{t+1}}$.

- Note use of misspecified model.
- Thus

$$f(x_t, y_t, \mathbb{E}_t[g(x_{t+1}, y_{t+1}, x_t, y_t]) \equiv -x_t + \mathbb{E}_t \left[g(x_{t+1}, y_{t+1}, x_t, y_t)\right]$$
$$= -q_t + \mathbb{E}_t \left[\widetilde{M}_{t+1} \times \left((1 - \overline{\delta})q_{t+1} + r_{t+1}^k\right)\right]$$

Auxiliary statistical model

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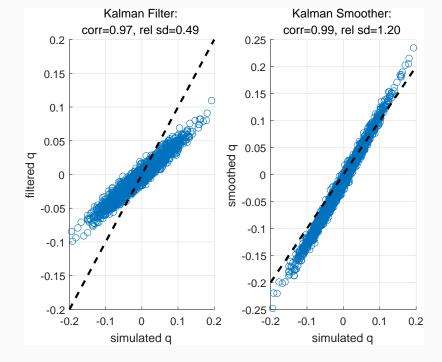
• Using that, in equilibrium, $1 + r_t^f \equiv \mathbb{E}_t[M_{t+1}]^{-1}$, we can re-write previous expression as an VAR approximation:

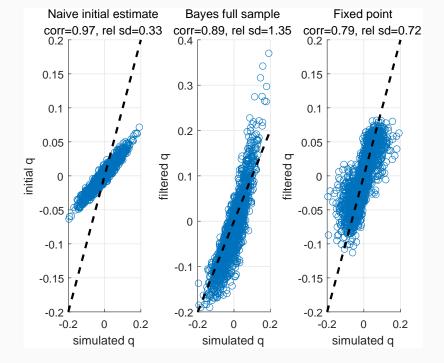
$$f(x_t^{(d)}, y_t, \widehat{\mathbb{E}}_t^{(d)}[g(x_{t+1}^{(d-1)}, y_{t+1}, x_t^{(d-1)}, y_t]) \approx \\ -x_t^{(d-1)} + (1 - \bar{\delta}) \Sigma_{1,4}^{(d)} + \Sigma_{1,2}^{(d)} \\ -(1 + r_t^f)^{-1} \times ((1 - \bar{\delta})e_4 + e_1) (\mu^{(d)} + A^{(d)} X_t^{(d-1)})$$

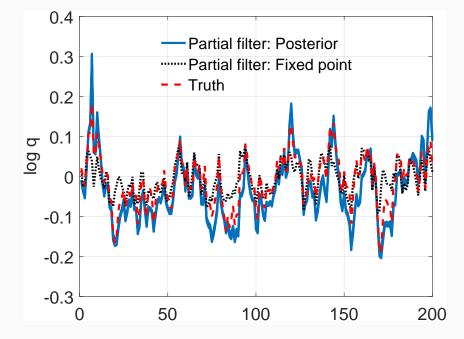
- We already are setting up notation for Gibbs sampler.
- For comparison purposes, we will also run:
 - 1. A naive guess for *q* computed as $q_t = \frac{(1-\bar{\delta})+mpk_{t+1}}{1+r_t^f}$.
 - 2. A Kalman filter and smoother.

VAR

- VAR(1) in $\xi_t^{(d)} \equiv [M_{t+1}, r_{t+1}^k, r_{t+1}^f, \widehat{q}_t^{(d)}]$:
 - 1. Fixed point: We solve for a fixed point in the 2,000-dimensional vector $\hat{q}^{(fp), T}$ and the VAR parameters $\hat{\mu}^{(fp)}, A^{(fp)}$, and $\Sigma^{(fp)}$.
 - 2. Gibbs sampler:
 - 2.1 Set $\xi_0^{(d)} = [M_0, r_0^k, r_0^f, \bar{q}]$
 - 2.2 We sample
 $$\begin{split} & \boldsymbol{\Sigma}^{(d)} \text{ and } \boldsymbol{\beta}^{(d)} = [\operatorname{vec}(\boldsymbol{A}^{(d)})', (\boldsymbol{\mu}^{(d)})']' \\ & \text{from} \\ & \boldsymbol{\Sigma}^{(d)} | \boldsymbol{\xi}^{(d), T} \sim \mathcal{IW}(T-1, \widehat{\boldsymbol{\Sigma}}^{(d)}_{OLS} \times T) \\ & \text{and} \\ & \boldsymbol{\beta}^{(d)} | \boldsymbol{\xi}^{(d), T}, \boldsymbol{\Sigma}^{(d)} \sim \mathcal{N}(\widehat{\boldsymbol{\beta}}^{(d)}_{OLS}, \boldsymbol{\Sigma}^{(d)} \otimes ((\overline{\boldsymbol{\xi}}^{(d-1)})'\overline{\boldsymbol{\xi}}^{(d-1)})^{-1}), \text{ conditioning on} \\ & \boldsymbol{\xi}^{(d)}_0 = [M_0, r_0^k, r_0^f, \overline{q}]. \end{split}$$
 - 2.3 We use a flat prior and define $\widehat{\Sigma}_{OLS}^{(d)} \times T$ as the OLS sum of squared residuals and $\widehat{\beta}_{OLS}^{(d)}$ the OLS estimator/MLE of the coefficients. $\overline{\xi}$ is a $T \times 5$ matrix with rows $\overline{\xi}_t = [\xi_{t-1}, 1]$.
 - 2.4 We solve optimality condition for $\hat{q}_t^{(d)}$.







Budd (2012), 'Labor Relations - Striking a Balance,' 4th ed.

A popular framework for thinking about labor law is to consider a **pendulum that** can range from strong bargaining power for labor ... to strong bargaining power for companies

We stress the role that changes in

- 1. statutory labor law (including executive orders),
- 2. case law (courts and NLRB), and
- 3. political climate

have on business cycles, income shares, and asset prices.

Model

- RBC model with search and matching frictions. (Andolfatto, 1996; Merz, 1995; Shimer, 2010)
 - Household with a continuum of members. Members are either employed or unemployed.
 - Household insures members against idiosyncratic employment risk.
 - Competitive firms that choose recruiting intensity.
 - Government.
 - Complete markets.
- Bargaining power subject to persistent redistribution shocks.

Households

• Recursive problem of the head of household:

$$V(a, n_{-1}) = \max_{a', n, c} \frac{c^{1-\sigma} (1 + (\sigma - 1)\gamma n_{-1})^{\sigma} - 1}{1 - \sigma} + \beta \mathbb{E}[V(a', n)]$$

with

$$c \equiv c_e n_{-1} + c_u (1 - n_{-1})$$

• Budget constraint:

$$c + \mathbb{E}[m' * a'] = (1 - \tau_n)wn_{-1} + T + a$$

with stochastic discount factor *m*.

• Law of motion of employment:

$$n = (1 - x)n_{-1} + f(\theta)(1 - n_{-1}),$$

with job finding rate $f(\theta) = \xi \theta^{\eta}$.

Firms

• Firm produces output y using effective capital uk_{-1} and production workers $(1 - \nu)n_{-1}$:

$$y = \left(\alpha^{\frac{1}{\varepsilon}}(uk_{-1})^{1-\frac{1}{\varepsilon}} + (1-\alpha)^{\frac{1}{\varepsilon}}(z(1-\nu)n_{-1})^{1-\frac{1}{\varepsilon}}\right)^{\frac{\varepsilon}{\varepsilon-1}},$$

Fraction ν workers devoted to recruiting activities.

• Laws of motion for employment and capital:

$$n = n_{-1}(\nu\mu(\theta) + 1 - x)$$

$$k = (1 - \delta(u))k_{-1} + I\left(1 - \frac{1}{2}\kappa\left(\frac{I}{k_{-1}} - \tilde{\delta}\right)^2\right)$$

where $\mu(\theta) = f(\theta)/\theta$ is hiring probability per recruiter.

• Firm value:

$$J(n_{-1},k_{-1}) = \max_{n,k,\nu} (1-\tau_k)(y-wn_{-1}) - I + \tau_k \delta(\bar{u})k_{-1} + \mathbb{E}[m' * J(n,k)]$$

Wage determination

- Generalized Nash bargaining between firms and households.
 - Workers have bargaining power ϕ .
 - Exogenous shifts in ϕ capture political shocks to bargaining process (Binmore *et al.*, 1986).
 - Other bargaining protocols? (Hall and Milgrom, 2008).
- Equilibrium wage solves

$$w = \arg \max_{\tilde{w}} \tilde{V}_n(\tilde{w})^{\phi} \tilde{J}_n(\tilde{w})^{1-\phi},$$

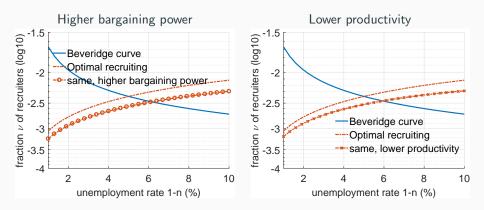
where \tilde{V}_n and \tilde{J}_n are marginal values of employment for households and firms given an arbitrary wage \tilde{w} .

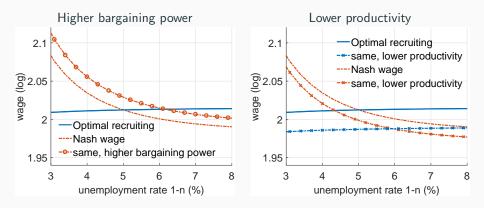
• Equilibrium wage along the balanced growth path:

$$ar{w} = ar{\phi} imes (1 + ar{ heta}) \overline{m{mpl}} + (1 - ar{\phi}) imes rac{\sigma}{1 - au_n} \left(rac{\gamma ar{m{c}}}{1 + (\sigma - 1) \gamma ar{m{n}}}
ight).$$

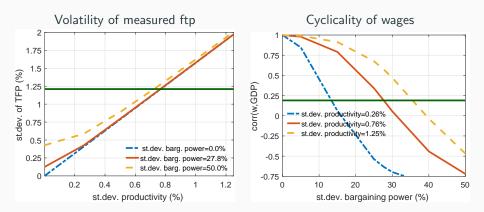
Equilibrium

- Government.
- Standard competitive equilibrium definition.
- Market clearing y = c + I.
- Aggregate capital and employment follow their law of motion.
- Two exogenous AR(1) shocks:
 - 1. Labor productivity z_t .
 - 2. Bargaining power ln $\frac{\phi_t}{1-\phi_t}$:
 - Baseline: half-life shocks of 8.5 years ≈ average control of presidency/house/senate after WWII.
 - Middle-run: half-life shocks of 20 years ≈ medium-term in Comín and Gertler (2006).
 - 2.3 Short run: half-life of 3.5 months.
 - 2.4 Long-run: new steady state.

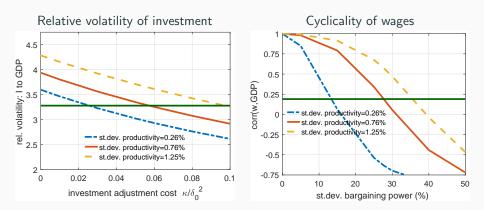




Identification (III)



Identification (IV)



Moment matching (I)

- Solve using pruned 3rd-order approximation (Andreasen et al., 2017).
- Select β , δ_0 , α , and τ_k to match moments from corporate non-financial business sector:
 - 1. 31.2% gross capital share.
 - 2. 12.7% gross depreciation share.
 - 3. 29.9% share of taxes in net surplus.
 - 4. 2.3 annual K/Y ratio.
- Match labor market statistics following Shimer (2010).
- Parametrized productivity and bargaining power process to match:
 - 1. 1.6% annual labor productivity growth.
 - 2. Volatility of measured z given persistence $0.95^{1/3}$.
 - 3. Cyclicality of wages.
 - 4. Relative standard deviation of investment I relative to Y.

Moment matching (II)

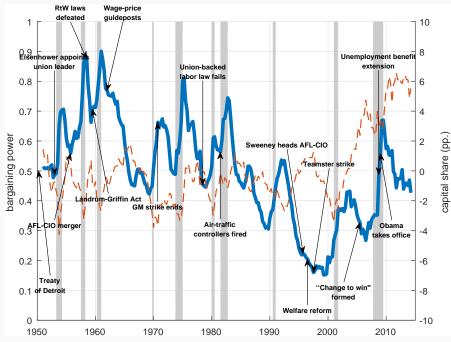
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Implementation of the partial filter

- Bargaining power enters only wage-setting.
- Wage-setting equation implies:

$$\begin{split} e^{\ln \frac{\phi_t}{1-\phi_t}} \left(\left(mpl_t \left(1 + \frac{1-x}{\mu(\theta_t)} \right) - w_t \right) \\ &- (1-x - f_t(\theta_t)) e^{\kappa_\phi + (\rho_\phi - 1)\ln \frac{\phi_t}{1-\phi_t} + \frac{1}{2}\omega_\phi^2} \left(\text{Cov}_t[\circ] + \frac{mpl_t}{\mu(\theta_t)} \right) \right) \\ &= w_t - \frac{1}{1-\tau_n} \left(\frac{c_t}{1+(\sigma-1)\gamma n_{t-1}} \right) \gamma \sigma, \quad \text{where} \\ &\text{Cov}_t[\circ] = \text{Cov}_t \left[\ln \frac{\phi_t}{1-\phi_t}, m_{t+1} \left(mpl_{t+1} \left(1 + \frac{1-x}{\mu(\theta_{t+1})} \right) - w_{t+1} \right) \right]. \end{split}$$

• Given Cov_t , solve for $\ln \frac{\phi_t}{1-\phi_t}$. Iterate in Gibbs-Sampler.



Concluding remarks

- We have other implementations.
- For example: sticky leverage of Gomes, Jermann, and Schmid (2016).
- However, many things to do:
 - 1. Embedding the model in an RBC model could aid in the calibration.
 - 2. Can we use machine-learning tools to improve the covariance/expectations computation?
 - 3. Heterogeneous agent model.
 - 4. Small sample results.
 - 5. Role for state smoothing due to estimation uncertainty/approximation error.