

Sovereign spreads, currency crises, and fundamentals: A non-linear analysis

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1 The issue

- Growing market size of emerging countries' bonds after the debt crisis of mid-eighties.
- Sovereign spreads: excess return of emerging markets debt instruments over comparable US Treasury bonds interest rate.
- Understanding spreads dynamic behaviour important for portfolio allocation and debt management...

⇒ large literature on fundamental determinants of sovereign spreads: Edwards (1984, 1986), Fernández Arias (1996), Kamin and Kleist (1999), Jostova (2006).

BUT Ferrucci (2003): non-fundamental factors play a more important role → over financial and currency crisis periods relationship between spreads and fundamentals breaks down → is this evidence of market inefficiency?

2 This paper

- THEORETICAL CONTRIBUTION: using Boschi and Goenka (2007) show that relationship between sovereign spreads and fundamentals weakens as pressure exerted by investors in foreign exchange markets increases.
- METHODOLOGICAL CONTRIBUTION: show how to bring to data a class of “second generation” models of financial/currency crises using a Markov-switching VAR framework \Rightarrow extension to a MULTIVARIATE analysis of Jeanne and Masson (2000).

3 The model

- The representative international investor maximizes expected utility over time.

$$U_t = E_t \left\{ \sum_{s=t}^{\infty} \beta^{s-t} u(C_s) \right\} \quad (1)$$

- E_t = expectation operator;
- $\beta \in (0, 1)$ = constant subjective time-preference factor;
- $u(\cdot)$ = period utility function.

- Period-by-period budget constraint:

$$\begin{aligned}
& B_{s+1}^{*f} + \sum_{m=1}^M x_{s+1}^{*m} B_s^{*m} + \sum_{m=1}^M x_{s+1}^m \frac{B_s^m}{e_s^m} = \\
& (1 + r_s^{*f}) B_s^{*f} + \sum_{m=1}^M x_s^{*m} (I_s^{*m} + B_s^{*m}) \\
& + \sum_{m=1}^M x_s^m \left(\frac{I_s^m + B_s^m}{e_s^m} \right) - C_s \quad (2)
\end{aligned}$$

- B_s^{*f} = risk free asset purchased at time s
- x_{s+1}^{*m}, x_{s+1}^m = fractions of country m 's security held
- B_s^{*m}, B_s^m = market value of country m 's dollar- and local currency debt

- r_s^{*f} = risk free interest rate

- I_s^m = dividend/coupon of country m 's security

- C_s = consumption.

- Maximizing (1) subject to (2) with respect to x_{s+1}^{*m} , x_{s+1}^m , and B_{s+1}^{*f} , gives the Euler equations at every date s for each asset.

- Manipulate the Euler equations; approximate through a Taylor expansion;

- We obtain:

$$E_t(r_{t+1}^{*m}) - r_{t+1}^{*f} \approx (1 + r_{t+1}^{*f})\beta \left(\frac{-C_t u''(C_t)}{u'(C_t)} \right) Cov \left\{ \frac{C_{t+1}}{C_t}, r_{t+1}^{*m} \right\}. \quad (3)$$

If we assume a decreasing relative risk aversion then a negative relationship between risk premium and aggregate consumption is established.

- Using a log-linear approximation to the budget constraint:

$$c_t - E_{t-1}c_t = (E_t - E_{t-1}) \sum_{j=0}^{\infty} \eta^j r_{t+j}^{*w} \quad (4)$$

- This provides the needed link between the risk premium and the probability of devaluation of country m' currency.

- If $\pi_{t+1}^m \uparrow$, $\implies E_t(r_{t+1}^{*m}) \downarrow \implies E_t r_{t+1}^{*w} \downarrow$.
 - $c_t - E_{t-1}c_t \downarrow \implies$ pure wealth effect.
 - Finally $\rho_{t+1}^m \uparrow$ through increased risk aversion.
- Build on a second generation model of currency crises in the spirit of Jeanne (1997).
 - Emerging economy will take any decisions about either maintaining the peg or not by minimizing the following loss function:

$$L_{t+1} = (u_{t+1})^2 + (\Delta d_{t+1})^2 + \delta \Gamma_{t+1} \quad (5)$$

- u_{t+1} = unemployment rate.

- Δd_{t+1} = growth in government real debt proportional to GDP.
- δ = dummy variable equal to 1 if devaluation occurs and 0 if the peg is maintained.
- Γ_{t+1} = exogenous cost of devaluing.
- Net benefit of the peg: difference between the *devaluation* and *fixed* loss function:

$$\bar{V}_{t+1} = L_{t+1}^d - L_{t+1}^f \quad (6)$$

- The decision rule of the policymaker is derived optimally:

$$\phi_{t+1} = E_t(V_{t+1}) \quad (7)$$

- The international investor formulates rational expectations:

$$\pi_{t+1} = \Pr [\bar{V}_{t+1} < 0]$$

- Thus:

$$\pi_{t+1} = \Pr [\epsilon_{t+1} < \alpha\pi_{t+1} + \varphi\rho_{t+1} - \phi_{t+1}] \quad (8)$$

Under certain conditions eq. (8) may have multiple solutions/equilibria.

4 Empirical illustration

Theory shows a complex interdependence between probability of currency crisis (π_{t+1}) and risk premium (ρ_{t+1}).

This relationship can be linearly approximated to obtain a Markov-switching VAR:

$$y_t = \nu(s_t) + \Pi x_t + \varepsilon_t \quad (9)$$

- $y_t = \begin{bmatrix} \pi_t \\ \rho_t \end{bmatrix};$

- $\nu(s_t)$ = intercepts allowed to switch between regimes;

- s_t = UNOBSERVABLE current regime/state \rightarrow follows hidden 2-state Markov chain stochastic process.

- \mathbf{x}_t = vector of relevant economic fundamentals.

I estimate a reduced-form, dynamic generalization of (9):

$$y_t = \nu(s_t) + \sum_{i=1}^p A_i y_{t-i} + \sum_{j=0}^q \Pi_j x_{t-j} + \varepsilon_t, \quad s_t = 1, \dots, n \quad (10)$$

\Rightarrow MSI(2)-VARX(p, q) model.

- Maximum likelihood estimation through *Expectation Maximization* (EM) algorithm.

4.1 Data

- Brazilian economy interesting case → three main periods of financial distress: 1994, 1998, 2001/2003.
- Proxy risk premium, ρ_t , using the Emerging Markets Bond Index + (EMBI+) developed by JP Morgan.
- Term “currency crisis” a strong speculative pressure on country’s currency \Rightarrow probability of devaluation, π_t , proxied by an index of the pressure exerted by participants on the foreign exchange market: Exchange Market Pressure Index (EMPI):

$$empi_t = -\frac{\sigma_e}{\sigma_{RES}} \Delta RES_t + \Delta e_t + \frac{\sigma_e}{\sigma_{id}} \Delta id_t$$

- RES_t = foreign exchange reserves;

- e_t = nominal exchange rate;
 - id_t = interest rate differential;
 - σ_i = standard deviation.
- Exogenous determinants: unemployment, u_t , government deficit, Δd_t , change in trade balance, Δtb_t , real exchange rate, $reer_t$, risk-free real interest rate, r_t^{*f} , rate of change of foreign exchange reserves, $\Delta rest_t$, rate of inflation, Δp_t , current account balance, ca_t , the rate of change of a broad money aggregate, Δm_t , spread between bank lending and deposit rates, ld_t , rate of change of domestic credit, $\Delta cred_t$, change in local stock market index, Δeq_t .

4.2 Results

Estimate a MSI(2)-VARX(1,1) over the period 1994:1-2005:12.

Table 4. MSI(2)-VARX(1,1) estimation results

	spread	empi
v(1)	3.84 (1.078)	0.618 (0.139)
v(2)	5.636 (1.201)	0.836 (0.152)
spread (-1)	0.752 (0.036)	-0.023 (0.005)
empi (-1)	-1.184 (0.480)	0.147 (0.063)
Δu	0.285 (0.212)	-0.018 (0.028)
Δd	25.006 (12.372)	1.036 (1.621)
Δtb	-22.917 (21.725)	0.219 (2.862)
reer	-8.922 (1.470)	-1.484 (0.194)
rf	-1.775 (30.899)	-1.172 (4.046)
Δp	-27.182 (5.975)	-1.020 (0.788)
Δip	-0.044 (0.043)	-0.010 (0.006)
Δca	24.572 (15.226)	1.412 (1.990)
Δm	10.086 (4.961)	0.341 (0.655)
ld	-0.659 (0.271)	-0.131 (0.035)
Δcred	11.902 (3.807)	0.822 (0.495)
Δres	-12.081 (2.892)	-3.191 (0.382)
Δeq	-0.002 (0.000)	-0.000 (0.000)
trend	-0.005 (0.003)	-0.001 (0.000)
σ	0.883	0.117
AIC	2.014 [2.105]	
HQ	2.361 [2.418]	
SC	2.867 [2.876]	
Log-lik	-101.992 [-112.481]	
LR linearity test statistic	20.977	
DAVIES	[0.0006]**	

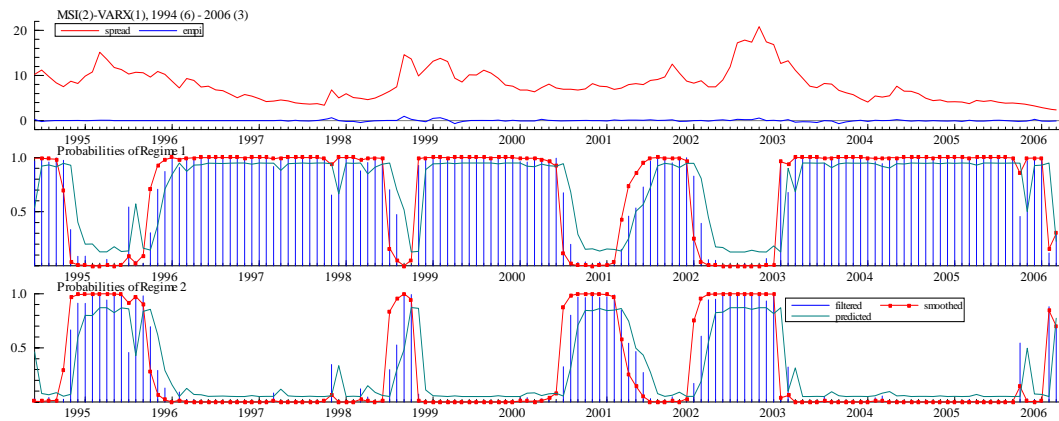
Notes: Coefficients in bold (italics) are significant at the 5% (10%) significance level. Standard errors are in parentheses. For the information criteria and the Log-likelihood, the value in square brackets is the corresponding statistics of the linear model.

Estimated matrix of transition probabilities shows two fairly persistent regimes do exist:

$$\hat{P} = \begin{pmatrix} 0.9489 & 0.0511 \\ 0.1278 & 0.8722 \end{pmatrix}$$

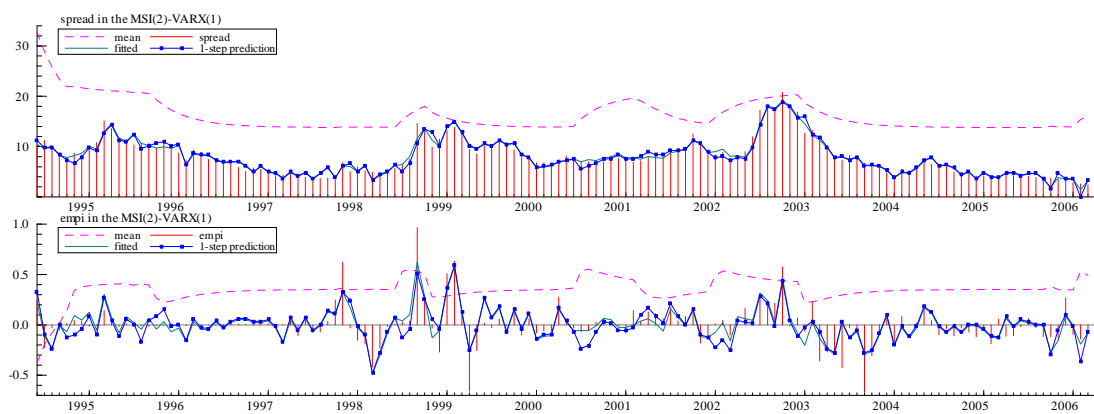
- Regime 1 corresponds to “tranquil” periods; Regimes 2 represents times of high economic tension → corresponds Mexican devaluation of 1994-1995, Russian default of 1998, Argentinean turmoil of 2001-2002, and pressures on the Brazilian currency surrounding the Presidential elections of 2002.
- Most of fundamental variables (8 out of 13) are statistically significant in either one of the two equations of the model, and have the right signs.
- The model performs better than a linear, i.e. one state, VAR(1) with same specification.

Figure 1: Regime probabilities



- Figure 5 shows the remarkable performance of the MS-VAR model by reporting the fitted and actual values of *spread* and *empi*.

Figure 2: MS(2)-VAR(1) model: fitted and actual values of Brazil's spreads and empi



5 Conclusion

- Earlier studies highlight that over crisis periods the relationship between sovereign spreads and macro-economic fundamentals breaks down, suggesting market inefficiency.
- This paper, by contrast, shows that a non-linear relationship between emerging markets spreads and fundamentals can be the outcome of investors' rational maximizing behaviour as they take into account the probability of domestic currency devaluations.
- Tests this hypothesis by developing an empirical version of the theoretical framework based on a Markov-switching regime multivariate model.
- Model outperforms its linear counterpart as well as univariate specifications, both linear and non-linear, of sovereign spreads.