

# **Shocking Markets: European Stock Markets and the ECB's Monetary Policy Surprises**

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## Motivation

- Measuring the consequences of unexpected monetary policy decisions on financial markets
- Of interest for central bankers (transmission process of interest rate setting) and financial market participants (portfolio decisions)
- Previous Evidence: Vast literature for the US perspective, little evidence for the European financial markets
- Our contribution to the literature:
  - (a) Applying the question to the ECB's monetary policy and the euro area stock markets
  - (b) Developing a methodology to extract monetary policy shocks
  - (c) Implementing an approach which controls for endogeneity of the variables

## Research Questions

- Is there a significant reaction of the stock markets following monetary policy shocks?
- How do the results fit in the existing international evidence?
- Does monetary impulses spread uniformly across major EMU stock markets?
- Does the ECB successfully communicates their monetary policy?

## Methodology

- A simple OLS is not suitable because of endogeneity
- A basic model for interaction of monetary policy and stock prices

$$\Delta i_t = \beta \Delta s_t + \gamma z_t + \epsilon_t \quad (1)$$

$$\Delta s_t = \alpha \Delta i_t + z_t + \eta_t \quad (2)$$

with

$\Delta i_t$  = change in interest rate

$\Delta s_t$  = stock returns

$z_t$  = vector of exogenous variables

$\epsilon_t$  = monetary policy shock

$\eta_t$  = asset price shock

## Methodology

Technique to handle endogeneity: Identification through heteroskedasticity approach (Rigobon and Sack (2004))

- Idea:  $\alpha$  can be identified without bias on days when the volatility of the monetary policy shock ( $\sigma_\epsilon$ ) increases
- Procedure: Watch the change in co-movement of interest rates and asset prices on days of higher compared to average variance of the monetary policy shock ( $\sigma_\epsilon$ )
- Define two subsamples
  - ◇  $F$  : Days of higher  $\sigma_\epsilon$  (ECB Governing Council meetings)
  - ◇  $F^*$ : Days of average  $\sigma_\epsilon$  (day preceding Council meetings)

## Methodology

- Assumptions on variance of the shocks
  - ◇  $\sigma_{\epsilon}^F > \sigma_{\epsilon}^{F^*}$ : Heteroskedasticity in variance of monetary policy shock
  - ◇  $\sigma_{\eta}^F = \sigma_{\eta}^{F^*}$ ,  $\sigma_z^F = \sigma_z^{F^*}$ : Homoskedasticity in variances of asset price and common shocks
- Identify  $\alpha$  out of the co-movement of interest rates and asset prices
  1. Solve Eq. (1) and (2)
  2. Estimate covariance matrix for each subsample  $F^*$  and  $F$
  3. Subtract matrix  $F^*$  from matrix  $F$  under the above stated assumptions
  4. The estimator  $\alpha$  can be identified as an element of the subtracted covariance matrix

## Data I

- Sample runs from January 1, 1999 to February 28, 2007
- Selected stock markets
  - ◇ The four largest national stock markets and their indices (German *DAX* 30, the French *CAC* 40, the Spanish *IBEX* 35 and the Italian *MIB* 30)
  - ◇ The aggregate European stock market (i.e. *Euro Stoxx* 50)
- Selected interest rate proxy
  - ◇ One month EURIBOR

## Data II

### Identifying monetary policy shocks

- Changes in expectation of market participants as proxy for monetary policy shocks
  1. Derived through heavily traded futures and swaps (EONIA swaps and EURIBOR futures)
    - ◇ Shocks extracted by principal components analyses (PC)
    - ◇ Shock if rate exceeds mean of the series by twice its standard deviation
  2. Derived through surveys of market economists (Reuters, Bloomberg, OBCE)
    - ◇ Shock if less than half of the respondents expected the move in advance
  3. Derived through a combination of both PC and survey definitions (COM)

# Empirical Results

Table 1: Shocks Extracted by Different Definitions

Date of Press Meeting	Shock Definition		
	PC	SURVEY	COM
08.04.1999	x		x
07.10.1999	x		x
03.02.2000		x	x
27.04.2000	x	x	x
08.06.2000	x		x
20.07.2000	x		x
05.10.2000	x	x	x
04.01.2001	x		x
29.03.2001		x	
11.04.2001		x	x
10.05.2001	x	x	x
30.08.2001		x	x
17.09.2001	x	x	x
11.10.2001	x	x	x
10.10.2002	x		x
07.11.2002	x		x
06.03.2003	x	x	x
08.05.2003		x	
01.12.2005		x	x
02.03.2006		x	
06.04.2006	x	x	x
05.10.2006		x	
07.12.2006		x	x
No. of shocks	14	16	19

## Empirical Results

Table 2: Stock Returns after Monetary Policy Shocks, 1 January 1999 - 28 February 2007

Shock Definition	PC	SURVEY	COM
Point Estimator	$\hat{\alpha}_{PC}$	$\hat{\alpha}_{SUR}$	$\hat{\alpha}_{COM}$
Euro Stoxx 50	-8.40 (-2.53)**	-9.30 (-2.30)**	-7.66 (-2.38)**
DAX 30	-7.78 (-2.10)*	-7.70 (-1.79)*	-7.01 (-2.03)*
CAC 40	-6.98 (-2.25)**	-7.19 (-1.85)*	-6.33 (-2.00)*
IBEX 35	-6.26 (-2.02)*	-6.48 (-1.77)*	-5.68 (-1.91)*
MIB 30	-4.32 (-1.31)	-3.61 (-0.97)	-4.16 (-1.38)

Notes: *PC* indicates shocks derived through the principal components analysis. *SURVEY* states the shocks found from the survey results. *COM* denotes extracted shocks as combination of the survey and the principal components definition. t-values are in brackets. The one month EURIBOR is applied as proxy for the policy rate. \*, \*\* denotes statistical significance at the 10 percent and 5 percent level, respectively.

## Summary and conclusion

- Findings for Europe indicate a negative and significant response of European stock returns (1.56 to 2.32 percent) to shocks (unexpected 0.25 percent rise) induced by the ECB
- Results correspond with the evidence for the US (Rigobon and Sack (2004), Bernanke (2005))
- Homogenous reaction of national stock markets to unexpected monetary policy decisions
- Number of shocks suggests a high predictability of the ECB's monetary policy

## Methodology: Supplements

- Solve the reduced form of Eq. (1) and (2) and estimate the covariance for both subsamples

$$\Omega_s = \frac{1}{(1 - \alpha\beta)^2} \begin{bmatrix} \sigma_\epsilon^s + \beta^2 \sigma_\eta^s + (\beta + \gamma)^2 \sigma_z^s & \alpha \sigma_\epsilon^s + \beta \sigma_\eta^s + (\beta + \gamma)(1 + \alpha\gamma) \sigma_z^s \\ \cdot & \alpha^2 \sigma_\epsilon^s + \sigma_\eta^s + (1 + \alpha\gamma)^2 \sigma_z^s \end{bmatrix} \quad (3)$$

with  $s = \{F, F^*\}$

- Assuming  $\alpha$ ,  $\beta$  and  $\gamma$  to be stable across the subsample, we can subtract the covariance matrices

$$\Delta\Omega = \Omega^F - \Omega^{F^*} = \frac{\sigma_\epsilon^F - \sigma_\epsilon^{F^*}}{(1 - \alpha\beta)^2} \begin{bmatrix} 1 & \alpha \\ \alpha & \alpha^2 \end{bmatrix} \quad (4)$$

and identify  $\hat{\alpha}$  by

$$\hat{\alpha} = \frac{\Delta\hat{\Omega}_{1,2}}{\Delta\hat{\Omega}_{1,1}} \quad (5)$$